

# A Generalization of Teicher's Poisson Tail Inequality

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## Abstract

[Teicher \[1955\]](#) proved that the probability a Poisson distribution with mean  $k$  takes on a value of  $k$  or less is monotonically decreasing in  $k$ . I extend this inequality by proving that the probability a Poisson distribution with mean  $zk$  takes on a value of  $k$  or less is monotonically decreasing for  $z \geq 1$ .

Let  $X_\lambda$  be a Poisson distributed random variable with mean  $\lambda$ . This note concerns the monotonicity of  $P(X_{zk} \leq k)$  in  $k$ . [Teicher \[1955\]](#) proves the following [see also: [Adell and Jodra, 2005](#)]:

$$(1) \quad P(X_k \leq k) > P(X_{(k+1)} \leq k+1)$$

**Proposition.**  $P(X_{zk} \leq k) > P(X_{z(k+1)} \leq k)$  for  $z \geq 1$ .

*Proof.* Let  $G_k(z) \equiv P(X_{zk} \leq k) - P(X_{z(k+1)} \leq k)$ . The proposition is equivalent to  $G_k(z) > 0$  for  $z \geq 1$ . Note that  $G_k(1)$  by [Adell, Jodra \(2005\)](#).

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Furthermore,  $\lim_{z \rightarrow \infty} G_k(z) = 0$ . Note that  $G_k(z)$  can be rewritten by the following steps:

$$(2) \quad G_k(z) = P(X_{zk} \leq k) - P(X_{z(k+1)} \leq k+1) = \\ P(X_{zk} \leq k+1) - P(X_{z(k+1)} \leq k+1) - P(X_{zk} = k+1)$$

By the Gamma-Poisson relationship, the lower tail of  $X_\lambda$  can be written as the upper-tail of a random variable with distribution  $\Gamma(k+1, 1)$ . Thus,  $P(X_\lambda \leq k) = P(Y > \lambda) = \frac{\int_\lambda^\infty u^k e^{-u} du}{k!}$ . Using this relationship, 2 can be written:

$$(3) \quad G_k(z) = \frac{\int_{zk}^\infty u^{(k+1)} e^{-u} du}{(k+1)!} - \frac{\int_{z(k+1)}^\infty u^{(k+1)} e^{-u} du}{(k+1)!} - \frac{(zk)^{k+1} e^{-(zk)}}{(k+1)!}$$

The first two terms have the same integrand but different bounds. The expression can be re-written as:

$$(4) \quad G_k(z) = \frac{\int_{zk}^{z(k+1)} u^{(k+1)} e^{-u} du}{(k+1)!} - \frac{(zk)^{k+1} e^{-(zk)}}{(k+1)!}$$

The derivative of  $(k+1)!G_k(z)$  with respect to  $z$  is:

$$(5) \quad (k+1)!G'_k(z) = (k+1)(z(k+1))^{(k+1)} e^{-z(k+1)} - (k)(zk)^{k+1} e^{-zk} \\ - k(k+1)(zk)^k e^{-zk} + k(zk)^{k+1} e^{-zk}$$

$$(6) \quad = (k+1)(z(k+1))^{(k+1)} e^{-z(k+1)} - k(k+1)(zk)^k e^{-zk}$$

Thus  $G'_k(z) > 0$  iff:

$$(7) \quad \left(\frac{k+1}{k}\right)^{(k+1)} > \frac{1}{z} e^z$$

For  $z = 1$  this is  $\left(\frac{k+1}{k}\right)^{(k+1)} > e$  which is true for all  $k > 0$ . Thus,  $G'_k(1) > 0$ . This together with [1](#) imply that  $G'_k(z)$  is initially positive and increasing on  $z \in [1, \infty)$ . Furthermore, there is a single stationary point since the equation  $\left(\frac{k+1}{k}\right)^{(k+1)} = \frac{1}{z}e^z$  has only one solution for  $z \geq 1$  which is given by the lower branch of the Lambert-W function:

$$(8) \quad z^* = -W\left(-\left(\frac{k}{k+1}\right)^{(k+1)}\right)$$

Thus,  $G_k(z)$  is strictly positive at  $z = 1$ , increases for  $z \leq z^*$  and then decreases for  $z \geq z^*$  approaching 0. This implies that  $G_k(z) > 0$  for  $z \geq 1$ .

□

## References

J. A. Adell and P. Jodra. The Median of the Poisson Distribution. *Metrika*, 61(3):337–346, June 2005. ISSN 0026-1335, 1435-926X. doi: 10.1007/s001840400350. URL <http://link.springer.com/article/10.1007/s001840400350>.

Henry Teicher. An inequality on poisson probabilities. *The Annals of Mathematical Statistics*, 26(1):147–149, 1955.